Asymmetric Shocks Inside Future EMU

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Abstract

In this paper, we focus on the issue of shocks asymmetries in the future European Monetary Union. Relying on the technique of codependence developed by Gouriéroux and Peaucelle [1989, 1993], we propose a new method which, unlike the analysis of Bayoumi and Eichengreen [1992], does not require any arbitrary restrictions. As a whole, our results support the idea of a two-speed Europe. (JEL Classification: C22, F33, F42)

I. Introduction

For a couple of years, the project of forming a European Monetary Union (EMU) has been extensively discussed. Both the timing and the exact composition of the future EMU have been particularly debated. Quite recently, it has been argued that a complete union with 12 or 15 countries was a non
viable arrangement. Indeed, even in the academic circles, a lot of authors support the idea of a “two-speed Europe” which favors the formation of a restricted monetary union with, for instance, Germany and her small neighbors followed by a later integration of some other countries having reached a satisfactory degree of economic convergence.

On the theoretical side, the plans for EMU have led to a renewal of the old Keynesian theory of optimum currency areas initiated by Mundell [1961]. This theory emphasizes some criteria that can be used to assess if a particular geographic area may or may not constitute a currency area. In recent formal models (Bayoumi [1994], Ricci [1995]), it has been shown once more that the degree of asymmetry of real shocks affecting the constituting countries of the zone is a key determinant of the net benefits drawn from a currency union.

Numerous empirical studies have been conducted to measure the magnitude of asymmetric disturbances between European countries. As shocks are not directly observable, the main problem and consequently the main source of disagreement between authors remain either their identification or the measurement of the degree of asymmetry. Unsurprisingly, this empirical literature has made extensive use of econometric methods which significantly differ from each other.

In this note, we present a new econometric approach aimed at assessing the degree of real shocks asymmetry. The method is based on the concept of codependence developed by Gouriéroux and Peaucelle [1989]. Unlike the structural VAR approach adopted by Bayoumi and Eichengreen [1992] which is considered as the reference paper on the topics, it allows to avoid the imposition of a priori restrictions on the parameters estimates. Section II quickly presents the main empirical studies. Section III provides a brief and intuitive presentation of the codependence technique. Section IV reports the results and emphasizes the limitations of the approach.

II. Empirical Evaluations of the Degree of Asymmetry

As a starting point, one can distinguish analyses relying on aggregated data such as GNP or total employment from those dealing with sectoral data. The latter are best illustrated by Bayoumi and Prasad [1995] as well as
by Helg et al. [1995].\(^1\) Most of the empirical analyses rely however on aggregated data.

As suggested by Vaubel [1976], the traditional approach is to analyze the real exchange variability (RER). This has been extensively applied to future EMU. For instance, using the variability of RER among six West German Länders as a yardstick, von Hagen and Neumann [1994] find empirical support to the idea of a two-speed Europe. In the same line of research, DeSerres and Lalonde [1995] attempt to distinguish between nominal and real shocks on the basis of the observed movements of RER. This approach faces, however, some criticisms. As pointed out by Eichengreen [1990], an important limitation is that it fails to distinguish between the effects of the shocks and the speed of adjustment to these shocks.

In order to account for this limitation, Bayoumi and Eichengreen [1992] use the structural vector autoregression (VAR) approach to isolating disturbances proposed by Blanchard and Quah [1989]. They identify, with some a priori restrictions on the effect on output and prices,\(^2\) supply and demand shocks across EC countries and US regions. Correlations between the computed disturbances provide some information about their symmetric nature, while the impulse response functions associated to the VAR give some insight into the speed of adjustment to these shocks.

Table 1 presents the main results of the literature in terms of the optimal composition of the future EMU. As a whole, this literature supports the idea of a two-speed Europe which discriminates between a core of countries hit by similar disturbances and a periphery, with some countries considered as intermediate cases.\(^3\) Nevertheless, there is no consensus about the exact composition of these groups.\(^4\)

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1. In particular, Helg et al. [1995] analyze the correlations between innovations of industrial production of 11 European countries on the basis of a cointegrated VAR. They distinguish between country-specific shocks, industry-specific shocks and local shocks.
2. These restrictions are however consistent with a simple aggregate supply and demand framework.
3. In a recent analysis, Rubin and Thygesen [1996] challenge this view.
4. Excepted for Belgium and the Netherlands which are generally expected to form a hard core with Germany.
III. Asymmetric Shocks and Codependence

The main default of the analysis of Bayoumi and Eichengreen [1992] is that the decomposition in supply and demand disturbances is based on arbitrary restrictions. Therefore, we propose an approach which does not require any a priori restriction on the data. We rely on the concept of codependence developed by Gouriéroux and Peaucelle [1989, 1993] and apply it to bilateral time series processes characterizing unemployment and real interest rates of several European economies.

Codependence techniques allow to account for long run relationships between stationary variables. The analysis is carried out in terms of dynamic multipliers and shock persistence in a multivariate moving average framework. We skip here the technical issues and, in particular, the estimation procedure. In order to get some insight into this approach, consider the following:

<table>
<thead>
<tr>
<th>Analysis</th>
<th>Method</th>
<th>Core</th>
<th>Periphery</th>
<th>Intermediate</th>
</tr>
</thead>
<tbody>
<tr>
<td>von Hagen and Neumann [1994]</td>
<td>Real Exchange Rate Variability</td>
<td>Ge, Au, Be, Ne, Fr</td>
<td></td>
<td>It, Dk, UK</td>
</tr>
<tr>
<td>DeSerres and Lalone [1995]</td>
<td>Real Exchange Rate Variability</td>
<td>Ge, Be, Ne</td>
<td>Sw, Fr, It</td>
<td>UK, Sp</td>
</tr>
<tr>
<td>Helg et al. [1995]</td>
<td>Sectoral Data / Cointegrated VAR</td>
<td>Ge, Be, Ne, Dk, Fr, UK</td>
<td>It, Sp</td>
<td>Gr, Ir, Pr</td>
</tr>
<tr>
<td>Bayoumi and Eichengreen [1992]</td>
<td>Structural VAR</td>
<td>Ge, Be, Ne, Dk, Fr</td>
<td></td>
<td>UK, It, Sp,</td>
</tr>
<tr>
<td>Artis and Zhang [1995]</td>
<td>Correlations between Cyclical Components</td>
<td>Ge, Be, Fr, Ne, Sp, Pr, lt</td>
<td>Ir</td>
<td>UK</td>
</tr>
</tbody>
</table>

5. Countries codes: Ge=Germany, Be=Belgium, Ne=The Netherlands, Fr=France, Au=Austria, Dk=Denmark, UK=United Kingdom, It=Italy, Sp=Spain, Pr=Portugal, Sw=Switzerland, Gr=Greece, Ir=Ireland.
6. The cyclical components are defined as deviations from a deterministic trend which is determined by a wide range of methods including the well-known Holdrick-Prescott decomposition.
7. The details can be found in Gouriéroux and Peaucelle [1989]. The behavior of the
lowing bivariate MA process of order $k$, for $k=2$:

$$
\begin{pmatrix}
Y_{1,t} \\
Y_{2,t}
\end{pmatrix} = \begin{pmatrix}
1, t \\
2, t
\end{pmatrix} + \begin{pmatrix}
21 & 22 \\
21 & 22
\end{pmatrix} \begin{pmatrix}
1, t-1 \\
2, t-1
\end{pmatrix} + \begin{pmatrix}
21 & 22 \\
21 & 22
\end{pmatrix} \begin{pmatrix}
1, t-2 \\
2, t-2
\end{pmatrix}
$$

(1)

Premultiplying by a vector $[1 \ -\alpha]$, this process reduces to a bivariate white noise. In this case, the order of reduction in the shocks persistence which is called the order of codependence (denoted $b$) is equal to 2. The existence of the vector $[1 \ -\alpha]$ implies that in the long run, the two variables $Y_{1,t}$ and $Y_{2,t}$ move towards an equilibrium defined by $Y_{1,t} - \alpha Y_{2,t} = \epsilon_{1,t} - \alpha \epsilon_{2,t}$ in which shocks adjust instantaneously.\(^8\)

In order to evaluate the degree of shocks asymmetry in such a bivariate framework, one can estimate the normalized codependence vector $[1 \ -\alpha]$ and measure how far it lies from the hypothetical vector $[1 \ -1]$ which is indicative of a symmetric (long run) evolution. In order to measure this distance, one may use, by analogy with the $R^2$, the following statistics $D$:

$$
D = [\cos (45^\circ \ - \arctg (\alpha))]^2.
$$

(2)

Furthermore, it is important to pay attention to the value of $k-b$ which is indicative of the speed of adjustment of shocks.\(^9\) Estimates as well as test statistics are a solution of a standard canonical correlations problem.

### IV. Results and Limitations

The procedure is applied to several European countries in a bivariate framework. Like in the Bayoumi and Eichengreen [1992] analysis, we choose Germany as the reference country. The analysis focuses here on the behavior of two macroeconomic variables which are expected to trace back

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\(^8\) Cases in which a linear combination does not yield a white noise but leads to a decrease in the VMA order are called non synchronous common cycles by Vahid and Engle [1993].

\(^9\) From a purely statistical point of view, a low degree of codependence may be indicative of a spuriously estimated codependence relationship due to problems of power.
the effects of aggregate real shocks, i.e. variations of unemployment rates and short run (ex-post) real interest rates. We use quarterly data. While the latter are found to be stationary on the estimation period (79Q2-94Q4) using the usual ADF unit root tests, the unemployment rates expressed in levels exhibit a stochastic trend. In this case, in order to prevent important power losses, it is advised (Beine and Hecq [1996]) to check for the existence of a cointegrating relationship between the variables in level. From the usual Johansen [1988] maximum likelihood procedure (not reported here to save place), it comes out that for the estimation period (70Q2-93Q4), the variables are not cointegrated.

Tables 2 and 3 present the results of the procedure. For each pair of countries, the tables report the value of \( k^{12} \), i.e. the order of the starting VMA.

### Table 2

**Test of Codependence – Variations of Unemployment Rates**

<table>
<thead>
<tr>
<th>70Q2 - 93Q4</th>
<th>( k )</th>
<th>( k - b )</th>
<th>Estimated Codependence Vector</th>
<th>( D )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ge/ Be(^{13})</td>
<td>6</td>
<td>2</td>
<td>([1 \ -0.634])</td>
<td>0.951</td>
</tr>
<tr>
<td>Ge/ Fr</td>
<td>4</td>
<td>2</td>
<td>([1 \ -2.95])</td>
<td>0.803</td>
</tr>
<tr>
<td>Ge/ UK</td>
<td>6</td>
<td>4</td>
<td>([1 \ -1.38])</td>
<td>0.974</td>
</tr>
<tr>
<td>Ge/ Dk</td>
<td>5</td>
<td>4</td>
<td>([1 \ 1.78])</td>
<td>0.072</td>
</tr>
<tr>
<td>Ge/ It</td>
<td>4</td>
<td>1</td>
<td>([1 \ -1.90])</td>
<td>0.912</td>
</tr>
<tr>
<td>Ge/ Ne</td>
<td>5</td>
<td>2</td>
<td>([1 \ -0.69])</td>
<td>0.966</td>
</tr>
<tr>
<td>Ge/ Ir</td>
<td>5</td>
<td>2</td>
<td>([1 \ -0.45])</td>
<td>0.872</td>
</tr>
<tr>
<td>Ge/ Sp</td>
<td>5</td>
<td>5</td>
<td>none</td>
<td>-</td>
</tr>
<tr>
<td>Ge/ Pr</td>
<td>4</td>
<td>4</td>
<td>none</td>
<td>-</td>
</tr>
</tbody>
</table>

10. Data sources: unemployment rates: standardized unemployment rates, OECD Business Sectoral Data Bank; prices: consumer price indexes, OECD main indicators; interest rates: 3-months interbank loans, mean values, National Bank of Belgium. The inflation rates were seasonally adjusted using exponential smoothing.
11. This result is also found by Kugler and Neusser [1993].
12. This value is chosen as to avoid underestimation of the VMA order which may induce in small samples important power losses (see on this point Beine and Hecq [1996]).
process, the value of $k - b$, i.e. the number of quarters the variables need to adjust towards the equilibrium relationship, the estimated vector and $D$, the measure of the distance from the hypothetical vector $[1 -1]$.

From these results, we can try to draw some conclusions in terms of composition of future EMU and compare them with the results reported in section II.

For a first group of countries constituted by Belgium, the Netherlands and Italy, our results do not produce any evidence of significant asymmetric disturbances with Germany. For these countries, the order of persistence of the innovations is quite small. For Belgium and the Netherlands especially, the estimated codependence vectors are very close from the hypothetical one $[1 -1]$ and the remaining persistence is relatively small.

For another group constituted by Spain and Portugal, our analysis concludes in favor of the existence of asymmetric disturbances. This is especially true for the analysis conducted for the changes in unemployment rates. To the extent that changes of unemployment rates are also a good proxy for global demand, this could suggest the presence of asymmetric demand shocks.

Finally, for countries like France, the UK and Denmark, the analysis concludes in favor of the existence of an asymmetric behavior in one of the two variables which have been investigated. This is the case for Denmark with

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respect to the variations of unemployment and for France and the UK for the analysis of real interest rates. On this basis, one may classify these countries as intermediate cases among the potential participants to EMU.

As a whole, these results confirm the idea of core-periphery distinction between the candidates to EMU which is in general supported by the empirical literature. Like in most of the previous studies, the small neighbors of Germany, i.e. Belgium, the Netherlands and to a less extent Austria are found to belong to the core. However, in terms of exact composition of the groups, our results are quite different from those of Bayoumi and Eichengreen [1992] which classify France as a core country and Italy as a peripheral one. This may confirm the view that as a whole, the empirical literature implementing the theory of optimum currency areas has been so far rather inconclusive in terms of optimal composition of future EMU (De Grauwe [1996]).

Finally, it is worth mentioning two important limitations of our approach. First, as illustrated by the analysis of the variations of unemployment, this method is not appropriate when the variables expressed in levels are cointegrated. Second, bivariate results may not be appropriate since EMU will be a multi-country currency area. Ideally, the method should investigate multi-country codependence to determine the exact size of the currency zone. These two points have been tackled in a recent analysis carried out by Rubin and Thygesen [1996] who rely on the notion of serial correlation common features developed by Engle and Kozicki [1993]. While interesting, this method yields results which have to be interpreted with caution.

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15. Note that in a recent analysis which tries to construct optimum currency areas indexes, Bayoumi and Eichengreen [1996] come to the opposite conclusion.
17. Basically, Rubin and Thygesen [1996] combine cointegration and codependence relationships applied to industrial production and prices in a multivariate set up. The concept of codependence is nevertheless rather different from ours. Actually, they look for serial correlation common features, i.e. for linear combinations which reduce the number of autoregressive terms. This in turn impedes a clear interpretation of the results in terms of shocks persistence which is offered by our approach. Indeed, if a linear combination of a starting VAR(p) process is not a white noise but just a VAR(p – q), both processes admit an infinite VMA representation.
References


Gouriéroux, Ch and I. Peaucelle [1993], “Séries codépendantes: application à l’hypothèse de parité du pouvoir d’achat,” in “Macroéconomie, Dévelop-
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